

## Solutions for Week13 Discussion Session

### 6.72

Given,  $Y_1$  and  $Y_2$  are independent and uniformly distributed over the interval  $(0,1)$ . So,  $F(y) = P(Y < y) = y; 0 < y < 1$ .

(a) We have to find the PDF of  $U_1 = \min(Y_1, Y_2)$ . We will use the CDF method. Hence,

$$\begin{aligned}F_{U_1}(u) &= P(U_1 < u) \\&= P(\min(Y_1, Y_2) < u) \\&= 1 - P(\min(Y_1, Y_2) > u) \\&= 1 - P(Y_1 > u)P(Y_2 > u); \text{ as they are independent} \\&= 1 - (1 - u)^2\end{aligned}$$

We have  $F_{U_1}(u) = 1 - (1 - u)^2; 0 < u < 1$

$$\implies \frac{dF_{U_1}(u)}{du} = f(u) = 2(1 - u); 0 < u < 1.$$

So, the pdf of  $U_1$  is:  $f(u) = 2(1 - u); 0 < u < 1$ . This is a  $\beta(1, 2)$  distribution.

(b) As  $f(u)$  is a beta distribution, hence, the  $E(U_1) = \frac{\alpha}{\alpha + \beta} = \frac{1}{3}$ .

$$\text{The } Var(U_1) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)} = \frac{1}{18}.$$

### 6.73

Given,  $Y_1$  and  $Y_2$  are independent and uniformly distributed over the interval  $(0,1)$ . So,  $F(y) = P(Y < y) = y; 0 < y < 1$ .

(a) We have to find the PDF of  $U_2 = \max(Y_1, Y_2)$ . We will use the CDF method. Hence,

$$\begin{aligned}F_{U_2}(u) &= P(U_2 < u) \\&= P(\max(Y_1, Y_2) < u) \\&= P(Y_1 < u)P(Y_2 < u); \text{ as they are independent} \\&= (u)^2\end{aligned}$$

We have  $F_{U_2}(u) = (u)^2; 0 < u < 1$

$$\implies \frac{dF_{U_2}(u)}{du} = f(u) = 2u; 0 < u < 1.$$

So, the pdf of  $U_2$  is:  $f(u) = 2u; 0 < u < 1$ . This is a  $\beta(2, 1)$  distribution.

(b) As  $f(u)$  is a beta distribution, hence, the  $E(U_2) = \frac{\alpha}{\alpha + \beta} = \frac{2}{3}$ .

$$\text{The } Var(U_2) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)} = \frac{1}{18}.$$

### 6.74

Given,  $Y_1, Y_2, \dots, Y_n$  are independent and uniformly distributed over the interval  $(0, \theta)$ . So, the cdf of  $Y_i$  are:  $F(y) = P(Y < y) = \frac{y}{\theta}; 0 < y < \theta$ .

(a) We have to find the PDF of  $Y_{(n)} = \max(Y_1, Y_2, \dots, Y_n)$ . We will use the CDF method. Hence,

$$\begin{aligned} F_{Y_{(n)}}(y) &= P(Y_{(n)} < y) \\ &= P(\max(Y_1, Y_2, \dots, Y_n) < y) \\ &= P(Y_1 < y)P(Y_2 < y)\dots P(Y_n < y); \text{ as they are independent} \\ &= \left(\frac{y}{\theta}\right)^n \end{aligned}$$

We have  $F_{Y_{(n)}}(y) = \left(\frac{y}{\theta}\right)^n; 0 < y < \theta$

(b) Now the pdf of  $Y_{(n)}$  will be:  $f_{Y_{(n)}}(y) = \frac{dF_{Y_{(n)}}(y)}{dy} = n\left(\frac{y}{\theta}\right)^{n-1}\frac{1}{\theta} = \frac{ny^{n-1}}{\theta^n}; 0 < y < \theta$ .

$$(c) E(Y_{(n)}) = \int_0^\theta y \frac{ny^{n-1}}{\theta^n} dy = \int_0^\theta \frac{ny^n}{\theta^n} dy = \frac{ny^{n+1}}{(n+1)\theta^n} \Big|_0^\theta = \frac{n\theta}{n+1}$$

To find the variance we first calculate  $E(Y_{(n)}^2)$ , then  $Var(Y_{(n)}) = E(Y_{(n)}^2) - [E(Y_{(n)})]^2$ .

$$E(Y_{(n)}^2) = \int_0^\theta y^2 \frac{ny^{n-1}}{\theta^n} dy = \int_0^\theta \frac{ny^{n+1}}{\theta^n} dy = \frac{ny^{n+2}}{(n+2)\theta^n} \Big|_0^\theta = \frac{n\theta^2}{n+2}$$

So,  $Var(Y_{(n)}) = \frac{n\theta^2}{(n+1)^2(n+2)}$ .

## 6.75

Given that the number of minutes that you need to wait for a bus is uniformly distributed on the interval  $[0,15]$ .

If you take the bus five times, that means we have five observations for waiting time in a bus stop  $Y_1, Y_2, Y_3, Y_4, Y_5$  respectively. We have to find the probability that longest wait is less than 10 minutes i.e.  $\max(Y_1, Y_2, Y_3, Y_4, Y_5) < 10$ .

Therefore,  $P(\max(Y_1, Y_2, Y_3, Y_4, Y_5) < 10) = P(Y_{(5)} < 10) = \left(\frac{10}{15}\right)^5$ ; using the distribution function of  $F_{Y_{(n)}}(y)$  from previous problem.

Hence, required probability is  $P(Y_{(5)} < 10) = (10/15)^5 = .1317$ .

## 6.80

Given that,  $Y_1, Y_2, \dots, Y_n$  be independent random variables, each with a beta distribution, with  $\alpha = \beta = 2$ .

So, the PDF of Y is  $f(y) = 6y(1-y); 0 < y < 1$  and the CDF is:  $F(y) = 3y^2 - 2y^3; 0 < y < 1$ .

(a) CDF of  $Y_{(n)}$  is:

$$\begin{aligned} F_{Y_{(n)}}(y) &= P(Y_{(n)} < y) \\ &= P(\max(Y_1, Y_2, \dots, Y_n) < y) \\ &= P(Y_1 < y)P(Y_2 < y)\dots P(Y_n < y); \text{ as they are independent} \\ &= (3y^2 - 2y^3)^n; 0 < y < 1 \end{aligned}$$

(b) The PDF is:  $f_{Y_{(n)}}(y) = \frac{dF_{Y_{(n)}}(y)}{dy} = 6ny(1-y)(3y^2 - 2y^3)^{n-1}; 0 < y < 1$

(c) We have  $n=2$ . So, from the above expression,  $f_{Y_{(2)}}(y) = 12y(1-y)(3y^2 - 2y^3); 0 < y < 1$

$$E(Y_{(2)}) = \int_0^1 y f_{Y_{(2)}}(y) dy = \int_0^1 12y^2(1-y)(3y^2 - 2y^3) dy = 0.6286$$